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FINANCIAL STRUCTURE AND ECONOMIC GROWTH

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CEPP WORKING PAPER NO. 06/05
June 2005

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Abstract

We address the issue of whether financial structure influences economic growth. Three competing views of financial structure exist in the literature: the bank-based, the market-based and the financial services view. Recent empirical studies examine their relevance by utilising panel and cross-section approaches. This paper for the first time ever utilises time series data and methods, along with the Dynamic Heterogeneous Panel approach, essentially on developing countries. We find significant cross-country heterogeneity in the dynamics of financial structure and economic growth, and conclude that it is invalid to pool data across our sample countries. We find significant effects of financial structure on real per capita output, which is in sharp contrast to some of the recent findings. Panel estimates, in most cases, do not correspond to country specific estimates, and hence may proffer incorrect inferences for several countries of the panel.

1. Introduction

Whether financial structure influences economic growth is a crucial policy issue. If one form of financial structure is more conducive to economic growth than another, then economic policy must take this into account. It is, therefore, hardly surprising that the distinction between bank-based and market-based financial systems, and their relative importance to economic growth, has been the focus of the relevant theoretical debate for over a century (Allen and Gale, 1999; Gerschenkron, 1962; Stiglitz, 1985). The debate is still very alive (see, for example, Levine, 2002), not least because resolving this issue undoubtedly improves the quality of economic policies.

The empirical literature on this issue attempts to examine whether one type of financial system better explains economic growth than another. However, these studies are not without their own problems. The studies that analyse the UK and the US as market-based systems versus Japan and Germany as bank-based systems (e.g. Hoshi et al, 1991; Mork and Nakkamura, 1999; Weinstein and Yafeh, 1998; Arestis et al., 2001) tend to show that financial structure matters. However, they are susceptible to the criticism that these countries historically share similar growth rates. Therefore, they may not form a suitable sample to investigate the relative contribution of one

financial system over another in the growth process (Goldsmith, 1969). Moreover, the results based on Japan, Germany, the UK and the US can only be used as a conjecture when it comes to economic policy for developing countries. Put it simply, the relationship between financial structure and economic growth remains unaddressed in the case of developing countries. We aim to fill this gap in this paper.

Panel and cross-section studies (Demirguc-Kunt and Levine, 1996; Levine, 2002 and 2003; Beck and Levine, 2002), find that financial structure is irrelevant to economic growth: neither the bank-based nor the market-based financial system can explain economic growth. Instead, it is the overall provision of financial services (banks and financial markets taken together) that are important. These multi-country studies are also subject to a number of concerns. Levine and Zervos (1996, p. 325) state that panel regressions mask important cross-country differences and suffer from 'measurement, statistical, and conceptual' problems. Quah (1993; see, also, Caseli et al., 1996) shows the difficulties associated with the lack of balanced growth paths across countries when pooling data. Pesaran and Smith (1995) point out the heterogeneity of coefficients across countries. Luintel and Khan (2004) show that panel estimates often do not correspond to country-specific estimates. Consequently, generalisations based on panel results may proffer incorrect inferences for several countries of the panel. In short, panel estimates may be misleading at country level; consequently their policy relevance may be seriously impaired.

This paper contributes to the empirical literature surrounding financial structure and economic growth in a number of ways. First, we collect time series data on financial structure for six countries, most of which are developing economies.¹ Data span for a minimum of 30 (South Korea) to a maximum of 39 years (Greece). Our sample of low- and middle-income countries with varied growth experiences (see Table 1), also addresses the concern raised by Goldsmith (1969). Second, we formally test whether the data of our sample countries can be pooled. Both time series and dynamic heterogeneous panel methods, which do not impose any cross-country restrictions on parameters and adjustment dynamics, are applied. To our knowledge, this is the first ever study of this kind which estimates the long-run relationship between financial

¹ The six countries are: Greece, India, South Korea, the Philippines, South Africa and Taiwan.

structure and economic growth using time series and dynamic heterogeneous panel methods, and utilising data from developing countries. Johansen's (1988, 1991) multivariate vector auto-regression (VAR), a well-established method in time series econometrics, is utilized. Panel estimates are obtained following the heterogeneous panel estimators due to Larson et al. (2001). In this way, our results could be compared with those in the relevant literature (e.g. Levine, 2002; Beck and Levine, 2002), and we could formally test whether the panel estimates (parameters) correspond to the country specific estimates.

Our results are quite revealing. First, time series results show that for the majority of sample countries financial structure significantly explains economic growth. The results from the dynamic heterogeneous panels also confirm the significance of financial structure. These findings are in sharp contrast to existing results, which either depict financial structure as irrelevant (Levin, 2002 and Beck and Levine, 2002), or else that only bank-based financial systems are conducive to growth (Arestis, et al., 2001). Second, we find significant heterogeneity in cross-country parameters and adjustment dynamics; tests show that data cannot be pooled for these six countries. Tests also show that the panel parameters do not correspond to country specific estimates. Thus, our results uphold the assertion of Levine and Zervos (1996) that 'panel regressions mask important cross-country differences'. Third, our results are robust to estimation methods and stability tests. Overall, our findings indicate that the apparent failure of large cross-country studies to identify a significant effect of financial structure on economic growth may be due to their failure to account sufficiently for cross-country heterogeneity.

The rest of the paper is organised as follows. In the section that follows we briefly discuss the theoretical arguments and the empirical evidence surrounding financial structure and economic growth. Section 3 outlines model specification and the econometric methods; section 4 discusses the dataset; section 5 covers tests of poolability; section 6 presents and discusses the main empirical results; and section 7 summarises and concludes.

2. Financial Structure and Development

Theoretical Considerations

The relationship between financial structure and economic development can be examined on the basis of competing theories of financial structure. These are: the bank-based, the market-based and the financial services. We discuss them briefly in what follows. The bank-based theory emphasises the positive role of banks in development and growth, and, also, stresses the shortcomings of market-based financial systems. It argues that banks can finance development more effectively than markets in developing economies, and, in the case of state-owned banks, market failures can be overcome and allocation of savings can be undertaken strategically (Gerschenkron, 1962). Those banks that are unhampered by regulatory restrictions, can exploit economies of scale and scope in information gathering and processing (for more details on these aspects of bank-based systems, see Levine, 2002, and Beck and Levine, 2002). Indeed, bank-based financial systems are in a much better position than market-based systems to address agency problems and short-termism (Stiglitz, 1985; Singh, 1997). The bank-based view also stresses the shortcomings of market-based systems. The latter reveal information publicly, thereby reducing incentives for investors to seek and acquire information. Information asymmetries are thus accentuated, more so in market-based rather than in bank-based financial systems (Boyd and Prescott, 1986). Banks can ease distortions emanating from asymmetric information through forming long-run relationships with firms, and, through monitoring, contain moral hazard. As a result, bank-based arrangements can produce better improvement in resource allocation and corporate governance than market-based institutions (Stiglitz, 1985; Bhidé, 1993).

By contrast, the market-based theory highlights the advantages of well-functioning markets, and stresses the problems of bank-based financial systems. Big, liquid and well-functioning markets foster growth and profit incentives, enhance corporate governance and facilitate risk management (Levine, 2002, and Beck and Levine, 2002). The inherent inefficiencies of powerful banks are also stressed, for they “can stymie innovation by extracting informational rents and protecting firms with close bank-firm ties from competition may collude with firm managers against other creditors and impede efficient corporate governance” (Levine, 2002, p. 3). Market-based financial systems reduce the inherent inefficiencies associated with banks and

are, thus, better in enhancing economic development and growth. A related argument is that developed by Boyd and Smith (1998), who demonstrate through a model that allows for financial structure changes as countries go through different stages of development, that countries become more market-based as development proceeds. An issue of concern, identified by a recent World Bank (2001) study in the case of market-based financial systems in developing countries, is that of asymmetric information. It is argued that “the complexity of much of modern economic and business activity has greatly increased the variety of ways in which insiders can try to conceal firm performance. Although progress in technology, accounting, and legal practice has also improved the tools of detection, on balance the asymmetry of information between users and providers of funds has not been reduced as much in developing countries as it has in advanced economies – and indeed may have deteriorated” (p. 7).

The third theory, the financial services view (Merton and Bodie, 1995; Levine, 1997), is actually consistent with both the bank-based and the market-based views. Although it embraces both, it minimises their importance in the sense that the distinction between bank-based and market-based financial systems matters less than was previously thought; it is financial services themselves that are by far more important, than the form of their delivery (World Bank, 2001). In the financial services view, the issue is not the source of finance. It is rather the creation of an environment where financial services are soundly and efficiently provided. The emphasis is on the creation of better functioning banks and markets rather than on the type of financial structure. Quite simply, this theory suggests that it is neither banks nor markets that matter; it is both banks and markets. They are different components of the financial system; they do not compete, and as such ameliorate different costs, transaction and information, in the system (Boyd and Smith, 1998; Levine, 1997; Demirguc-Kunt and Levine, 2001). Under these circumstances, financial arrangements emerge to ameliorate market imperfections and provide financial services that are well placed to facilitate savings mobilisation and risk management, assess potential investment opportunities, exert corporate control, and enhance liquidity. Consequently, as Levine (2002) argues, “the financial services view places the analytical spotlight on how to

create better functioning banks and markets, and relegates the bank-based versus market-based debate to the shadows” (p. 3).²

Empirical Evidence

A number of studies have concentrated on comparisons that view Germany and Japan as bank-based systems, while the US and UK as market-based systems. These studies have employed rigorous country-specific measures of financial structure. Studies of Germany and Japan use measures of whether banks own shares or whether a company has a ‘main bank’ respectively (Hoshi et al., 1991; Mork and Nakkamura, 1999; Weinstein and Yafeh, 1998). These studies provide evidence that confirms the distinction between bank-based and market-based financial systems in the case of the countries considered. However, reassessment of the evidence on the benefits of the Japanese financial system in view of the economy’s poor performance in the 1990s has concluded against the beneficial effects of the bank-based nature of this system. Bank dependence can lead to a higher cost of funds for firms, since banks extract rent from their corporate customers (Weinstein and Yafeh, 1998). Studies of the US and the UK concentrate on the role of market takeovers as corporate control devices (Wenger and Kaserer, 1998; Levine, 1997), and conclude in favour of market-based financial systems. Goldsmith (1969), however, argues that such comparisons in the case of Germany and the UK for the period 1864-1914 does not contribute to the debate since “One cannot well claim that a superiority in the German financial structure was responsible for, or even contributed to, a more rapid growth of the German economy as a whole compared to the British economy in the half-century before World War I, since there was not significant difference in the rate of growth of the two economies” (p. 407). Our own study (Arestis et al., 2001) that provides evidence for the superiority of bank-based systems may be subjected to the same

² A special case of the financial services view is the law and finance view (La Porta et al, 1998; see, also, Levine, 1999). It maintains that the role of the legal system in creating a growth-promoting financial sector, with legal rights and enforcement mechanisms, facilitates both markets and intermediaries. It is, thereby, argued that this is by far a better way of studying financial systems rather than concentrating on bank-based or market-based systems. The World Bank (2001) view on the matter, based on “econometric results systematically points in one direction: far from impeding growth, better protection of the property rights of outside financiers favors financial market development and investment” (p. 8). Indeed, Rajan and Zingales (1998) argue that although countries with poor legal systems benefit from a bank-based system, better legal systems improve market-based systems, and as such the latter are preferable. While we recognise the importance of legal systems in a growth-promoting finance sector, we do not attempt to deal with this issue in this paper. It requires a study by itself and as such it is left for another occasion.

criticism. We may note in passing, though, that the implications for developing economies are evident, as argued in that paper.

Levine (2002) reinforces Goldsmith's (1969) argument when concluding that "financial structure did not matter much since the four countries have very similar long-run growth rates" (p. 4). Levine (op. cit.) addresses this problem by using a broad cross-country approach that allows treatment of financial system structure across many countries with different growth rates. The findings of this study support neither the bank-based nor the market-based views; they are, instead, supportive of the financial services view, that better-developed financial systems is what matters for economic growth. An earlier study by Demirguc-Kunt and Levine (1996), using data for forty-four industrial and developing countries for the period 1986 to 1993, had concluded that countries with well-developed market-based institutions also have well-developed bank-based institutions; and countries with weak market-based institutions also have weak bank-based institutions. Thereby supporting the view that the distinction between bank-based and market-based financial systems is of no consequence. However, Levine and Zevros (1998), employing cross-country regressions for a number of countries covering the period 1976 to 1993, conclude that market-based systems provide different services from bank-based systems. In particular, market-based systems enhance growth through the provision of liquidity, which enables investment to be less risky, so that companies can have access to capital through liquid equity issues (see, also, Atje and Jovanovic, 1993, and Harris, 1997). The World Bank (2001) provides a comprehensive summary of the available evidence, which also reaches similar conclusions. It argues strongly that the evidence should be interpreted as clearly suggesting that "both development of banking and of market finance help economic growth: each can complement the other" (p. 48). In what follows we attempt to tackle the problem alluded to by Goldsmith (1969) and others. We also deal with the concerns surrounding the panel and cross-country regressions referred to by, among others, Levine and Zervos (1996). Our usage of time-series and heterogeneous panel estimators to analyse a number of diverse countries, should go some way in tackling these concerns.

3. Specification and Econometric Methods

Specification

We specify a generalised Cobb-Douglas production function of the following form:

$$\log(Q/L)_t = a_0 + a_1 \log(K/L)_t + a_2 \log(\text{STR})_t \quad (1)$$

where Q is output, L is labour, K is capital and STR is financial structure (defined as the market capitalisation over bank lending; see, also, below). Higher STR means a system that is more of the market-based variety; while a lower STR means more of a bank-based system. In specification (1), financial structure directly accounts for Total Factor Productivity (TFP). In actual estimations we use per capita output (LYP) and per capita capital stock (LKP), since consistent time series on labour force do not exist for most of the sample countries. It is important to note that for the purposes of this study, we are interested in the significance or otherwise of the coefficient a_2 , rather than its sign. In either case a significant a_2 coefficient implies that financial structure matters; an insignificant a_2 coefficient implies that financial structure is of no consequence whatsoever.

It is common that cross-section studies use several other determinants of economic growth - the years of schooling (human capital), black market premiums, indicators of civil liberty, revolutions and coups, assassinations, bureaucratic efficiency, corruptions etc. However, data on these variables are usually obtained from periodic surveys and hence consistent time series are unavailable. Nevertheless, our specification (1) compares quite favourably with the ‘simple conditioning set’ specified by Levine (2002) and Beck and Levine (2002). They use initial levels of income and schooling as ‘simple conditioning set’ and examine the effect of financial structure on economic growth in panel/cross sectional framework, whereas we specify a generalised Cobb-Douglas production function.³

³ Barro and Lee (2000) and Cohen and Soto (2001) provide periodic data, five yearly and 10 yearly respectively, on educational attainment for several countries of the world. We thought of interpolating annual series on educational attainments for our sample countries from these periodic observations. However, the owners of the respective data sets advised us strongly against interpolation, on the grounds of unreliability. We, therefore, decided not to pursue this matter any further.

Econometric Methodology

Under the Johansen (1988) maximum likelihood (ML) approach a k -dimensional and p^{th} order vector (X) can be re-parameterised to a vector error-correction model (VECM):

$$\Delta X_t = \mu + \Gamma_1 \Delta X_{t-1} + \Gamma_2 \Delta X_{t-2} + \dots + \Gamma_{p-1} \Delta X_{t-p+1} + \Pi X_{t-p} + \varphi D_t + \varepsilon_t \quad (2)$$

In our analysis X_t [LYP, LKP, STR] is a 3×1 vector of first-order integrated [I(1)] variables; Γ_i are (3×3) short-run coefficient matrices; $\Pi_{(3 \times 3)}$ is a matrix of long-run (level) parameters; D_t captures the usual deterministic components; μ is a constant term and ε_t is a vector of Gaussian error. A co-integrated system, X_t , implies that: (i) $\Pi = \alpha_{(3 \times r)} \beta'_{(r \times 3)}$ is rank deficient, i.e. $r < k$ (r = number of distinct cointegrating vectors; $k = 3$); and (ii) $\{\alpha_{\perp} \Gamma \beta_{\perp}\}$ has full rank, $(k-r)$, where α_{\perp} and β_{\perp} are $(3 \times (3-r))$ orthogonal matrices to α and β . The rank of Π is tested by the well known Maximal Eigenvalue (λ -max) and Trace statistics (Johansen, 1988).

The Johansen method is a reduced-form dynamic system estimator, which addresses the issues of multi-cointegration and normalisation. A number of issues are important for the estimation of the VAR model. It is the time span of the data rather than the number of observations, which determines the power of cointegration tests (Campbell and Perron, 1991). Our data extend from a minimum of 30 (South Korea) to a maximum of 39 (Greece) years, which in our view, provides sufficient time length to capture the long-run relationship between LYP, LKP and STR. We specify the VAR lengths (p) such that the VAR residuals are rendered non-autocorrelated.⁴ A constant term is entered in the co-integrating space to allow for non-zero mean of the system variables. A trivariate VAR can exhibit two cointegrating vectors at the most. Pesaran and Shin (2002) suggest identification of multi-cointegration through the tests of over-identifying restrictions. We follow their approach of identification if multiple cointegrating vectors are found.

⁴ Johansen (1992) suggests that the lag length in the VAR should be specified whereby the VAR residuals are rendered uncorrelated. Selection of lag length based on information criteria may not be adequate to render the VAR residual uncorrelated (Cheung and Lai, 1993). Hence, we specify lag-length based on the test of serial correlation in VAR residuals.

4. Description of Data

Our sample consists of six countries, viz. Greece, India, South Korea, the Philippines, South Africa and Taiwan. Data on real Gross Domestic Product (GDP), real gross fixed Investment (I), Bank Lending (BLR, defined as total lending by deposit taking institutions) and total population are obtained from IMF CD-ROM (March 2002). Market capitalization (CLR, defined as total value of domestic equities listed in domestic stock exchanges) is obtained from Global Financial Data, Inc. and Standard and Poor's Emerging Markets database (2002). Data frequency, determined simply by data availability, is annual and the sample period is 1962-2000 for Greece, 1970-1999 for South Korea, 1966-1999 for India, 1969-1999 for the Philippines, 1965-1999 for South Africa and 1965-2000 for Taiwan. The heterogeneity in the sample period across countries is due to unavailability of data. A consistent series of total physical capital stock for the whole sample period is, unfortunately, not available. Therefore, we constructed it for each country in the sample from the respective real gross fixed investment series using the perpetual inventory method. Following Luintel and Khan (1999), amongst others, a depreciation rate of eight percent and the sample-average growth rate of real investment, are used to compute the initial capital stock. Following Levin (2002) and Beck and Levine (2002), financial structure is defined as the log of the market capitalisation over bank-lending. Thus, our measure of financial structure is akin to their measure of 'structure-size'.

Table 1 reports some descriptive statistics of our data set. It is obvious that our sample consists of countries with differing income levels and varied growth experience. In our sample Korea was the fastest growing economy (7.6 % growth of real per capita income per annum) and the Philippines the slowest (0.6% per capita income growth per annum). A striking feature, however, is that all but one (Taiwan) sample countries have evolved towards a more market-based system over the last thirty to forty years. Although the bank lending ratios have gone up for all sample countries during this period, the rise in capitalisation ratio is by far greater. The bank-lending ratio, on average, is 2.73 folds higher in the last five years of the sample compared to its level of first five year, but the capitalisation ratio has shot up by 7.60 folds during the same period. The low base may partly explain this huge rise in the capitalisation ratio. Nevertheless, financial structure, on average, has gone up by almost three folds in the intervening period. The last column of Table 1 shows that the financial systems of sample countries grew towards a more market-oriented system by an average annual

rate of 0.2 % for Taiwan (lowest) and 2.5% for Korea (highest). Figure 1 plots LYP for all sample countries.

All plots are normalised at 1995=1 for the ease of comparison across countries. In econometric modelling we do not use normalised data. It is apparent that LYP shows positive trend for all countries but at varying rate. Taiwan and South Korea show quite steep rise in real per capita output whereas plots for India and the Philippines appear relatively flat. Plots for Greece and South Africa are in between.

Figure 2 plots LKP. Again, the rate of capital accumulation appears quite high for Taiwan and South Korea, as their plots of LKP are pretty steep. Greece shows a rapid rate of capital accumulation prior to 1975, but it slows down thereafter. For the remaining countries the rate of accumulation process appears rather slow as depicted by the flatness of their plots.

Figure 3 plots the financial structure variable. Plots appear more volatile than those of LYP and LKP; a positive trend is also not easily discernable in some cases.⁵ Greek financial structure depicts big spikes during the 1970s and late 1990s. Taiwan and South Korea also show spikes during late 1980s but of a lesser magnitude. Overall, these plots show a gradual move towards a market oriented financial system.

5. Heterogeneity

Our sample consists of low- and middle-income countries, which represent different stage of development and economic structure. They also share significantly different growth experiences (Table 1). It is, therefore, interesting to formally test if it is valid to pool the data set of these countries. This is important not least because there is a growing concern about the panel and cross-section tests, in that they neglect heterogeneity.

Formal tests of the dynamic heterogeneity of financial structure and economic growth are conducted as follows. First, we estimate a series of p^{th} ($p=1,2,3$) order autoregressive and distributed lag models, ADL(P), conditioning LYP on LKP and STR, and test for the equality of parameters across sample countries. Second, we estimate ADL(P) on growth rates and perform tests of parameter equality.

⁵ This lack of apparent positive trend in some of these plots is mainly due to the big scaling on the vertical axis, which is required to accommodate the Greek series. Country-by-country plot shows a positive trend more clearly.

Chow-type F tests under the null of parameter equality across sample countries are reported in Table 2, where the tests reject the null under all specifications. Thus, the elasticity of LYP with respect to LKP and STR is heterogeneous across countries. Furthermore, as another measure of dynamic heterogeneity, we test for error variances homoskedasticity across groups. The LM-test of group-wise heteroskedasticity is reported in Table 2, which confirms that error variances across sample countries are significantly different and this also holds across all specifications. It follows that the elasticity of LYP with respect to LKP and STR, as well as the error dynamics across sample countries, are significantly heterogeneous. Consequently, the data set cannot be pooled. This raises concerns with respect to the validity of extant panel, and cross-sectional tests that do not allow for cross-country heterogeneity.

6. Empirical Results

In order to evaluate the time series properties of the data formally, we implement the univariate KPSS test (Kwiatkowski et al., 1992), which tests the null of stationarity.⁶ The results are reported in Table 3. LYP and LKP are non-stationary; tests reject the null of stationarity in all cases but one, this being the Philippines' LYP. The latter's trend stationarity is rejected but level stationarity is not. The financial structure variable also appears non-stationary in all but two cases. The exceptions are Greece and South Korea. The Greek financial structure appears level stationary but not trend stationary whereas the opposite holds for South Korea. We further examine the autocorrelation functions for these (three) suspects and found that they decay slowly which means they appear closer to I(1) series than to I(0). Hence, we treat them as I(1) in further modelling. All series appear unequivocally stationary in their first differences. Thus, the overall finding of the KPSS tests is that LYP, LKP and STR are I(1).

⁶ Kwiatkowski et al. (1992) show that these tests are more powerful than the usual DF/ADF tests. Recently, however, Caner and Kilian (2001) warn against these power gains, especially for high frequency data. Our data are low frequency.

Table 4 reports the Johansen rank tests and a range of VAR diagnostics obtained from the VECM. Trace tests show that LYP, LKP and STR are co-integrated and exhibit a single co-integrating rank (vector) for all sample countries. The λ -max statistic supports these findings except that the evidence of co-integration for India and the Philippines now appears marginal. Given the superiority of trace statistics over the maximal eigen-value statistics in testing the null of non-co-integration, we conclude that LYP, LKP and STR are co-integrated with single co-integrating vector for all of our sample countries.

For a valid normalisation and error-correction representation, the associated loading factors (α_s) must be negatively signed and significant. On this basis, we can normalise all countries on LYP; their associated loading factors are negatively signed and significant at 5% or better. Given the signs of loading factors and the existence of a single co-integrating vector, the parameters of our empirical model are uniquely identified. LM tests show absence of serial correlation in VAR residuals for all cases. The VAR residuals pass normality tests except for Taiwan. Thus, utilizing the VECM, we are able to identify a long-run output relationship based on per capita capital stock and the financial structure variable, which confirms the error-correcting behaviour when displaced from the long-run equilibrium.

Table 5 (section A) reports the estimated co-integrating vectors (long-run parameters). As expected the long-run elasticity of LYP with respect to LKP is positive and highly significant for all countries included in our sample. Typically, large panel studies, which do not account for cross-country parameter heterogeneity, estimate the contribution of capital stock within the range of 0.30 to 0.40. Our average point estimate (0.506) is on the higher side and country-specific parameters exhibit heterogeneity. Financial structure significantly affects per capita GDP for all but one country (the Philippines). The sign and the significance of parameters show that the market-based financial system appears conducive to Greece, India, South Korea and Taiwan, whereas the bank-based system appears better for South Africa in explaining the long-run per capita output. In the case of the Philippines, financial structure appears insignificant in explaining per capita output. Overall, we find significant

effects of financial structure on long-run per capita output in the majority of cases examined (five out of six countries).

In section B of Table 5 we report the panel, ‘between-dimension’, estimates of the elasticities of LYP with respect to LKP and STR. Larsson et al. (2001) discuss the computations of these panel estimates under the Johansen approach. Essentially, the ‘between-dimension’ panel parameters are the mean of the country-specific parameters. Our panel-based results corroborate our time series findings: financial structure appears significant for this panel of sample countries. This is in contrast to the findings of Levine (2002) and Beck and Levine (2002). It is important to note that the negatively signed large coefficient (-0.519) of South Africa alone is sufficient to turn the overall coefficient for the panel into negative (-0.008). All the existing panel tests suffer from this typical caveat – results of one or few countries dominate the whole panel – and one of the contributions of our results is that they bring this issue to focus.⁷ This further lends support to our preference to country-by-country (time series-based) results.

A note on the reconciliation of our results with the existing panel results is in order. The ‘between-dimension’ panel approach we use differs from the panel approach followed by Levin (2002) and Beck and Levine (2002). In fact, the cointegration-based ‘between-dimension’ panel test (which we implement) is a statistically superior test than that implemented by Levine (op. cit.) and Beck and Levine (op. cit.). This is because the ‘between-dimension’ approach allows for the cross-country heterogeneity, whereas the approach of Levin (2002) and Beck and Levine (2002) does not. This difference in empirical approach may explain the differences in the two sets of results. We reiterate that we attach more importance to the country specific time-series results, and to the cross-country heterogeneity results exhibited; this is precisely the focus of the paper. The fact that the data employed for the purposes of this paper cannot be pooled across our sample countries lends further support to the appropriateness of the time-series approach as opposed to the panel approach. The

⁷ Panel unit root tests, panel co-integration tests (dynamic heterogeneous or otherwise) and traditional (OLS- and/or IV- based) panel tests all suffer from this problem.

‘between dimension’ dynamic heterogeneous panel results are, therefore, reported merely to highlight the point that panel results may be misleading.

The magnitudes of the country-specific point estimates (elasticities) in Table 5 show a considerable degree of cross-country heterogeneity. From a policy perspective, it is extremely important to establish the degree of equivalence between the panel and the country specific estimates because national policies rely on such key parameters. We address this issue by formally testing if country-specific parameters are jointly equal to the corresponding panel estimates. This involves conducting a Wald or LR test for the restriction that each country-specific coefficient is equal to its panel counterpart and summing up the individual χ^2 statistics (Pesaran et al., 2000). Assuming that these tests are independent across countries, the sum of the individual χ^2 statistics tests for the null that country-specific coefficients are jointly equal to their respective panel estimates. The test statistic is $\chi^2(N)$ distributed, where N is the number of countries in the panel. The empirical test statistic significantly rejects the null that the financial structure variable exerts the same effect across countries.⁸ The cross-country heterogeneity in the parameters of financial structure is significantly different. Thus, our results show that financial structure significantly affects the level of output for most sample countries (five of the six), and the cross-country heterogeneity in estimated point elasticities is pervasive and statistically significant.

Stability

It is well known that structural shifts should be identified endogenously rather than exogenously (see, among others, Perron, 1997; Christiano, 1992; Quintos, 1995; Luintel, 2000). Hence, we evaluate the stability of estimated co-integrated vectors by following the recursive approach of Hansen and Johansen (1999). This approach essentially compares the recursively-computed ranks of the Π matrix with its full sample rank. A significant difference between them implies a structural shift in the cointegrating rank. Likewise, conditional on the identified ranks of Π , if sub-sample parameters significantly differ from those of the full sample, this signifies instability of the cointegrating parameters. The LR test for these hypotheses is asymptotically χ^2 , with $kr-r^2$ degrees of freedom. Tests are carried out in two settings: (i) allowing both

⁸ The χ^2 test statistic under the null of parameter equality is $\chi^2(6)=30.440$.

short-run and long-run parameters to vary (the Z-model); and (ii) short-run parameters are fixed and only long-run parameters are allowed to vary (the R-model).

We specify a base estimation window of the first 20 observations.⁹ Since sample sizes differ across countries, the period of stability tests is not uniform – the stability tests range between a minimum of 13 years (South Korea) to a maximum of 17 years (Greece). Figure 4 plots the normalised LR statistics that test rank stability using the R-model.¹⁰ All LR statistics are scaled by the 5% critical value; hence, values greater than unity imply rejection of the null of stability and vice versa. In these plots stability of rank, r , requires rejection of $r-1$ ranks.

Plots of the scaled LR statistics show that the null of non-cointegration ($H_0: r=0$) is clearly rejected for all sample countries. All plots that test the null of $r=0$ cross the critical threshold. The plots that test $H_0: r \leq 1$ are all below unity (i.e. less than the 5% critical value) except for a marginal break shown by India during 1994-95. Figure 5 plots the normalised LR statistics, which test for the stability of cointegrating parameters.

Plots pertaining to both Z- and R-models are reported. Co-integrating parameters are stable for all countries with only one exception. The Z- model shows parameter instability for South Africa prior to 1990; this is primarily due to the volatility of short-run parameters since the R-model shows parameter stability. Overall, our estimated co-integrating rank and parameters are remarkably stable.

7. Summary and Conclusions

We have focused in this paper on the important, but controversial, issue of whether financial structure matters in an economic system. We briefly reviewed the relevant theoretical and empirical literature before embarking on a time-series investigation of this issue, the first of this kind to be investigated. We further provided panel results based on dynamic heterogeneous panel estimator.

⁹ Hansen and Johansen (1999) specify an initial estimation window of 16 (monthly) observations.

¹⁰ The R-model is more suitable for testing the stability of cointegrating ranks and long-run parameters (Hansen and Johansen, 1999). Nonetheless, results from the Z-model appear broadly similar and hence are not reported (but are available from the authors upon request).

Our results clearly show that significant cross-country heterogeneity exists in financial structure and growth dynamics and it is invalid to pool data even for these six countries. This indicates that extant panel and/or cross-section studies of financial structure and economic growth, which pool several countries, may well have concealed important cross-country differences. We find a robust co-integrating relationship between output per capita, capital stock per capita and the financial structure. Financial structure exerts significant effects on the level of output per capita in all but one country (the Philippines). Furthermore, the magnitude of the long-run effects (cointegrating parameter) of financial structure on per capita output is extremely heterogeneous across countries. Tests reject the null of equality between the ‘between-dimension’ panel and country specific parameters vis-à-vis the financial structure variable. Thus, panel estimates do not appear to correspond to country specific estimates (parameters). The speed of adjustment to long-run disequilibria also differs significantly across countries. A comparison of our time series and panel results also reveals that a single country may sufficiently dominate the result for the whole panel. Consequently, panel results may provide deceptive results for most country estimates in the panel.

Our findings of a significant effect of financial structure on output levels are in sharp contrast to those of Levine (2002) and Beck and Levine (2002), amongst others. This contrast is maintained by both the empirical approaches (time series and the dynamic heterogeneous panel estimators) we have pursued in this study. We attribute this difference in the results to our empirical approach, which allows for cross-country heterogeneity in parameters and adjustment dynamics. It is, thus, possible that the apparent insignificant effect of financial structure on growth shown by extant panel tests may be due to their failure to address cross-country heterogeneity. The main policy message of our findings is that financial structure matters for economic growth.

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Table 1: Some Summary Statistics of Data

	Growth of Real Per Capita Income	Bank Lending Ratio (BLR)				Capitalisation Ratio (CLR)				Financial Structure (STR)			
		BLR ₀	BLR _T	BLR _μ	ΔBLR	CLR ₀	CLR _T	CLR _μ	ΔCLR	STR ₀	STR _T	STR _μ	ΔSTR
Greece	0.037	0.273	0.921	0.720	0.020	0.043	0.757	0.198	0.024	0.156	0.807	0.302	0.020
India	0.026	0.248	0.462	0.409	0.007	0.036	0.337	0.113	0.007	0.146	0.729	0.254	0.013
South Korea	0.076	0.393	0.703	0.520	0.017	0.066	0.368	0.224	0.024	0.167	0.516	0.401	0.025
Philippines	0.006	0.281	0.671	0.384	0.011	0.119	0.674	0.254	0.011	0.423	1.030	0.632	0.005
South Africa	0.015	0.900	1.431	1.026	0.019	0.876	1.549	1.043	0.022	0.976	1.086	1.010	0.004
Taiwan	0.064	0.346	1.867	0.978	0.045	0.183	1.033	0.417	0.020	0.550	0.554	0.373	0.002
Average Change	0.037	2.733				7.603				2.968			

BLR = total lending by deposit taking institutions/GDP; CLR = total value of domestic equities listed in domestic stock exchange/GDP; STR = log(CLR/BLR). Growth of Real Per Capita Income indicates the average annual growth rate (expressed as ratio) over the sample period. Subscripts 0 and T denote mean values of the first five years and the last five years of the sample for each country; subscript μ indicates the sample mean value. The average change (in the last row) is calculated as: $\sum_1^6 (X_t - X_0)/6$, where X denotes BLR, CLR and STR.

Table 2: Tests of heterogeneity of financial structure and growth dynamics across sample countries

	Specification: A			Specification: B		
	P=1	P=2	P=3	P=1	P=2	P=3
Parameter equality	17.352 ^a (5, 171)	19.299 ^a (7, 153)	12.083 ^a (10, 129)	9.389 ^a (5, 165)	8.577 ^a (7, 147)	5.833 ^a (10, 123)
LM Test	14.508 ^b	15.450 ^a	12.492 ^b	11.169 ^b	11.115 ^b	12.551 ^b

The specification A: $LYP = \lambda_0 + \sum_{i=1}^p \lambda_{1i} LYP_{t-i} + \sum_{i=1}^p \lambda_{2i} LKP_{t-i}^d + \sum_{i=1}^p \lambda_{3i} STR_{t-i} + \varepsilon_t$.

The specification B: $\Delta LYP = \theta_0 + \sum_{i=1}^p \theta_{1i} \Delta LYP_{t-i} + \sum_{i=1}^p \theta_{2i} \Delta LKP_{t-i} + \sum_{i=1}^p \theta_{3i} \Delta STR_{t-i} + \varepsilon_t$.

Equality of θ and λ are standard (Chow type) F-tests of parameter equality across the sample (six) countries. Numbers within parentheses, (.), are the degrees of freedom of F distribution. The 1% critical value for F(5, 125) is 3.17; the 5% critical value for F(5, 125) is 2.29. Lagrange Multiplier (LM) tests (see text) reject the null of homoskedastic error variances across the sample countries; they are $\chi^2(5)$ distributed. Superscripts 'a' and 'b' indicate rejection of the null at 1% and 5%. Variable definitions are: LYP = log of per capita real GDP; LKP = log of per capita real physical capital stock; and STR = log (CLR/BLR).

Table 3: KPSS unit root tests

Countries	LYP		LKP		STR	
	η_{μ}	ι_{μ}	η_{μ}	ι_{μ}	η_{μ}	ι_{μ}
Greece	1.256 ^a	0.2338 ^a	1.254 ^a	0.326 ^a	0.207	0.198 ^b
India	1.151 ^a	0.248 ^a	1.182 ^a	0.292 ^a	0.704 ^b	0.278 ^a
South Korea	1.079 ^a	0.132 ^c	1.086 ^a	0.134 ^c	0.642 ^b	0.078
Philippines	0.312	0.136 ^c	0.893 ^a	0.259 ^a	0.452 ^c	0.186 ^b
South Africa	0.994 ^a	0.266 ^a	0.866 ^a	0.319 ^b	0.505 ^b	0.116 ^c
Taiwan	1.393 ^a	0.204 ^b	1.379 ^a	0.291 ^a	0.368 ^c	0.194 ^b

The critical values for η_{μ} are 0.739, 0.463 and 0.347 at 1%, 5% and 10%; the respective critical values for ι_{μ} are 0.216, 0.146 and 0.119. η_{μ} and ι_{μ} respectively test the nulls of level and trend stationarity. In their first differences all series are stationary. The latter set of results is not reported to conserve space; they are available from the authors upon request. Superscripts a, b, and c indicate rejection of the null of stationarity at 1%, 5% and 10%, respectively. Variable definitions are: LYP = log of per capita real GDP; LKP = log of per capita real physical capital stock; and STR = log(CLR/BLR).

Table 4: Co-integration tests and VAR diagnostics between LYP, LKP and STR (Johansen Method)

	Trace Statistics, H_0 :			Maximal Eigenvalue, H_0 :			Loading Factor (α)	LM{3}	NOR	LAG
	r=0	r≤1	r≤2	r=0	r≤1	r≤2				
GRE	39.53 ^b (0.015)	15.63 (0.196)	4.19 (0.396)	23.89 ^b (0.027)	11.44 (0.228)	4.19 (0.396)	-0.401 ^a [0.078]	0.748	0.192	2
IND	34.53 ^b (0.057)	14.85 (0.241)	3.00 (0.589)	19.68 (0.113)	11.85 (0.201)	3.00 (0.588)	-0.919 ^a [0.248]	0.188	0.334	2
KOR	32.60 ^c (0.092)	10.03 (0.643)	2.98 (0.593)	22.58 ^b (0.043)	7.05 (0.669)	2.98 (0.592)	-0.404 [0.123]	0.453	0.128	3
PHL	36.24 ^b (0.037)	17.46 (0.117)	6.40 (0.168)	18.78 (0.148)	11.07 (0.256)	6.40 (0.167)	-0.652 ^b [0.296]	0.795	0.115	2
SAF	34.92 ^b (0.052)	11.79 (0.477)	2.36 (0.707)	23.13 ^b (0.036)	9.43 (0.402)	2.36 (0.706)	-0.185 ^a [0.046]	0.114	0.215	2
TWN	44.40 ^a (0.003)	17.86 (0.110)	7.22 (0.120)	26.53 ^a (0.010)	10.64 (0.290)	7.22 (0.120)	-0.683 ^a [0.138]	0.146	0.000	3

The country mnemonics are: GRE=Greece; IND=India; KOR= South Korea; PHL= the Philippines; SAF= South Africa; TWN=Taiwan. Figures within parenthesis (.) are p-values under the H_0 : r=0; r ≤ 1 and r ≤ 2. For the loading factors (α) figures within the brackets [.] are standard errors. LM{3} reports p-values of the third order LM test of the null of no serial correlation in VAR residuals. The column NOR reports p-values of Bera-Jarque normality tests of VAR residuals, $\chi^2(2)$ distributed. The column LAG reports the VAR lag lengths used. Superscript a, b, and c indicates significance at 1%, 5% and 10% respectively. GRE, IND and PHL do not require any dummy. KOR required impulse dummies for 1978 and 1998; Taiwan required an impulse dummy around first oil price shock (1970 and 1971), although non-normality is still evident. SAF required impulse dummies for 1981 and 1987. These impulse dummies are entered unrestricted in the VAR. Exclusion of these dummies does not change the results qualitatively except for the failure of the diagnostics (non-normality and/or autocorrelation). Superscripts a, b, and c indicate significance at 1%, 5% and 10%, respectively.

Table 5: Estimated Cointegrating parameters (Normalised on LYP)

	Section A: Country-by-Country time series parameters						Section B
	GRE	IND	KOR	PHL	SAF	TWN	Panel results
LKP	0.431 ^a [0.066]	0.594 ^a [0.043]	0.637 ^a [0.040]	0.273 ^a [0.031]	0.487 ^a [0.145]	0.615 ^a [0.035]	0.506 ^a [0.046]
STR	0.067 ^a [0.023]	0.054 ^a [0.012]	0.168 ^a [0.045]	0.007 [0.008]	-0.519 ^a [0.143]	0.176 ^a [0.062]	-0.008 ^b [0.004]

The country mnemonics are: GRE=Greece; IND=India; KOR=Korea; PHL= the Philippines; SAF= South Africa; TWN=Taiwan. Figures within parenthesis [.] are respective standard errors. Variable definitions are: LYP = log of per capita real GDP; LKP = log of per capita real physical capital stock; and STR = log(CLR/BLR). Superscripts a, b, and c indicate parameter significance at 1%, 5% and 10%, respectively. Section B reports dynamic heterogeneous between dimension panel results as discussed in the text.

Figure 1: GDP per capita (1995=1)

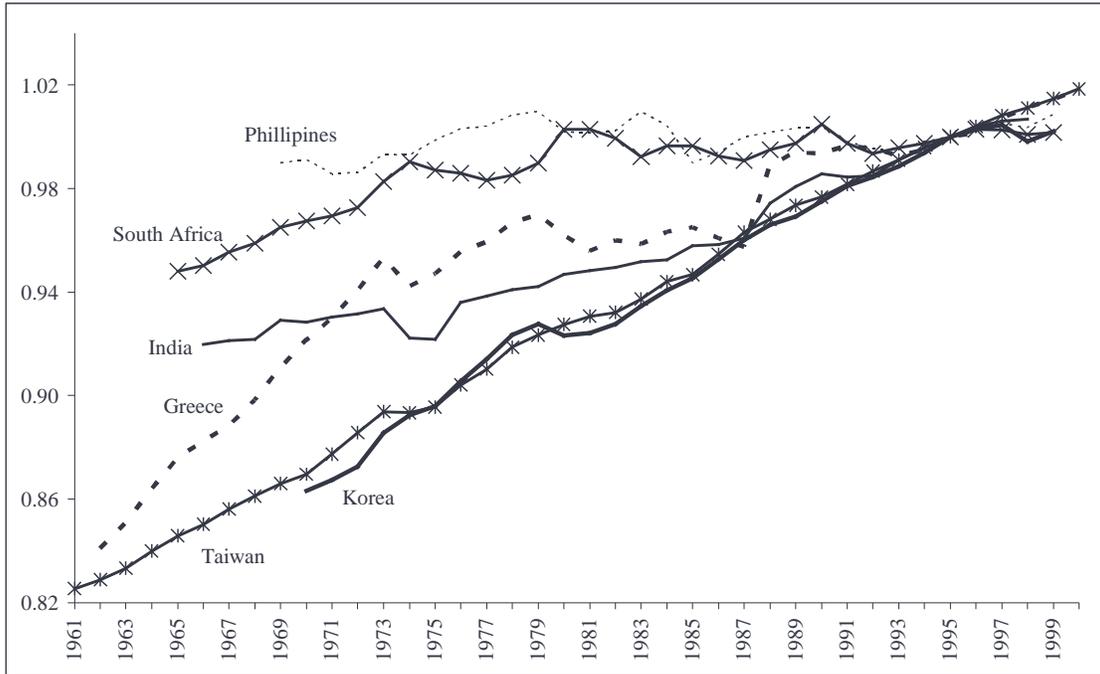


Figure 2: Per capita capital stock (1995=1)

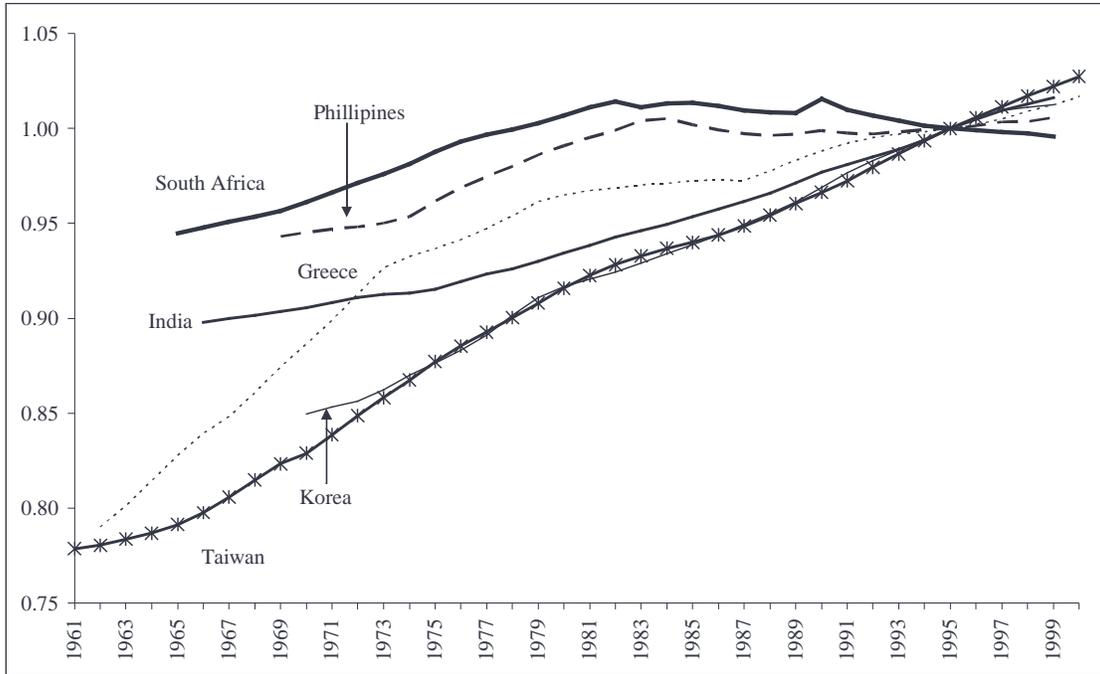


Figure 3: Financial structure variable (1995=1)

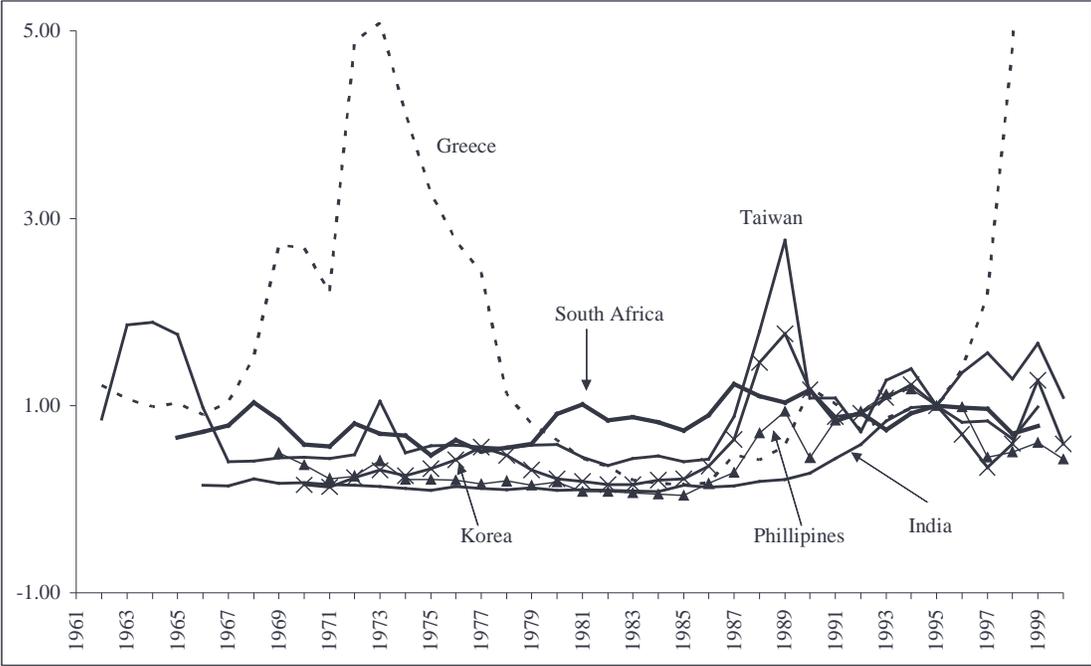


Figure 4: Plots of Scaled Recursive LR- Statistics (Rank Stability Tests: R Model)

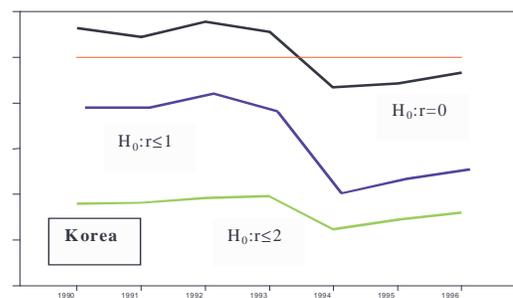
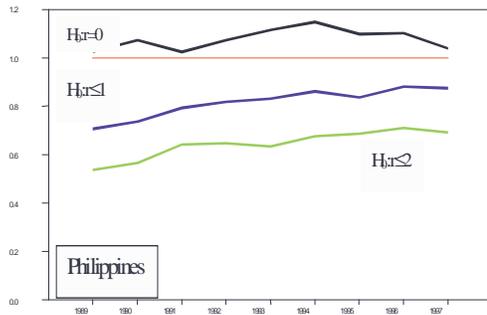
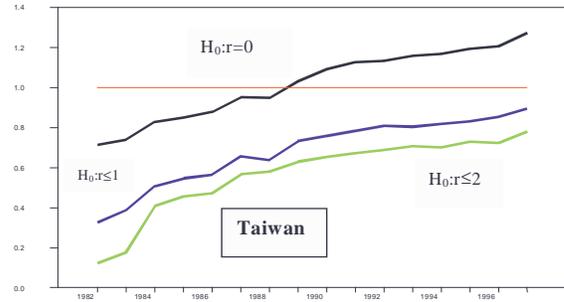
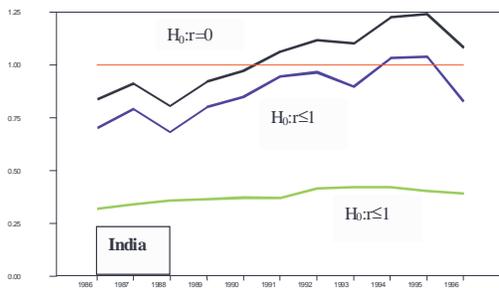
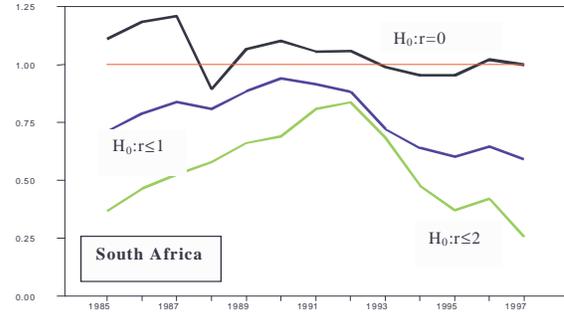
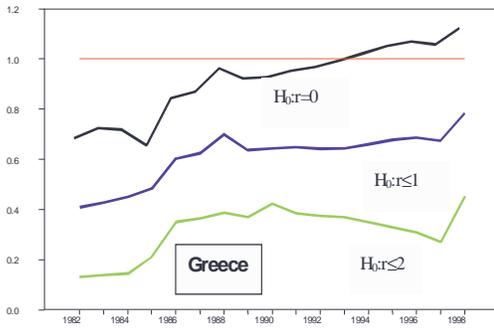


Figure 5: Parameter Stability Tests (Plots of Scaled Recursive LR- Statistics)

